## Lecture 2 Binomial and Poisson Probability Distributions

## **Binomial Probability Distribution**

- Consider a situation where there are only two possible outcomes (a Bernoulli trial)
  - ★ Example:
    - flipping a coin
      - 🛯 head or tail
    - rolling a dice
      - ☞ 6 or not 6 (i.e. 1, 2, 3, 4, 5)
  - $\star$  Label the probability of a success as p
    - the probability for a failure is then q = 1 p
- Suppose we have *N* trials (e.g. we flip a coin *N* times)
  - what is the probability to get *m* successes (= heads)?
- Consider tossing a coin twice. The possible outcomes are:
  - \* no heads:  $P(m=0) = q^2$
  - ★ one head: P(m = 1) = qp + pq (toss 1 is a tail, toss 2 is a head or toss 1 is head, toss 2 is a tail) = 2pq ← two outcomes because we don't care which of the tosses is a head
  - **\*** two heads:  $P(m = 2) = p^2$

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- \*  $P(0) + P(1) + P(2) = q^2 + 2pq + p^2 = (q + p)^2 = 1$
- We want the probability distribution function P(m, N, p) where: m = number of success (e.g. number of heads in a coin toss)
  - N = number of trials (e.g. number of coin tosses)
  - p =probability for a success (e.g. 0.5 for a head)

James Bernoulli (Jacob I) born in Basel, Switzerland Dec. 27, 1654-Aug. 16, 1705 He is one 8 mathematicians in the Bernoulli family (from Wikipedia).



• If we look at the three choices for the coin flip example, each term is of the form:

 $C_m p^m q^{N-m}$  m = 0, 1, 2, N = 2 for our example, q = 1 - p always!

- $\star$  coefficient  $C_m$  takes into account the number of ways an outcome can occur regardless of order
- ★ for m = 0 or 2 there is only one way for the outcome (both tosses give heads or tails):  $C_0 = C_2 = 1$
- ★ for m = 1 (one head, two tosses) there are two ways that this can occur:  $C_1 = 2$ .
- Binomial coefficients: number of ways of taking N things m at time

$$C_{N,m} = \binom{N}{m} = \frac{N!}{m!(N-m)!}$$

- ★  $0! = 1! = 1, 2! = 1 \cdot 2 = 2, 3! = 1 \cdot 2 \cdot 3 = 6, m! = 1 \cdot 2 \cdot 3 \cdots m$
- $\star$  Order of things is not important
  - e.g. 2 tosses, one head case (m = 1)
    - we don't care if toss 1 produced the head or if toss 2 produced the head
- ★ Unordered groups such as our example are called *combinations*
- ★ Ordered arrangements are called *permutations*
- ★ For *N* distinguishable objects, if we want to group them *m* at a time, the number of permutations:  $P_{N} = \frac{N!}{m}$

$$P_{N,m} = \frac{1}{(N-m)!}$$

- example: If we tossed a coin twice (N = 2), there are two ways for getting one head (m = 1)
- example: Suppose we have 3 balls, one white, one red, and one blue.
  - Number of possible pairs we could have, keeping track of order is 6 (rw, wr, rb, br, wb, bw):

$$P_{3,2} = \frac{3!}{(3-2)!} = 6$$

• If order is not important (rw = wr), then the binomial formula gives

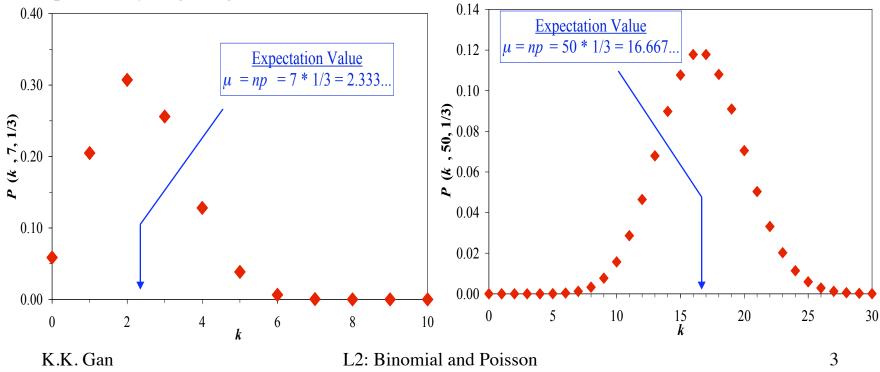
$$C_{3,2} = \frac{5!}{2!(3-2)!} = 3$$
 number of two-color combinations

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• Binomial distribution: the probability of *m* success out of *N* trials:

$$P(m,N,p) = C_{N,m} p^m q^{N-m} = \binom{N}{m} p^m q^{N-m} = \frac{N!}{m!(N-m)!} p^m q^{N-m}$$

- p is probability of a success and q = 1 p is probability of a failure
- Consider a game where the player bats 4 times:
  - ★ probability of  $0/4 = (0.67)^4 = 20\%$
  - ★ probability of  $1/4 = [4!/(3!1!)](0.33)^1(0.67)^3 = 40\%$
  - ★ probability of  $2/4 = [4!/(2!2!)](0.33)^2(0.67)^2 = 29\%$
  - ★ probability of  $3/4 = [4!/(1!3!)](0.33)^3(0.67)^1 = 10\%$
  - ★ probability of  $4/4 = [4!/(0!4!)](0.33)^4(0.67)^0 = 1\%$
  - ★ probability of getting at least one hit = 1 P(0) = 0.8



• To show that the binomial distribution is properly normalized, use Binomial Theorem:

$$(a+b)^{k} = \sum_{l=0}^{k} {\binom{k}{l}} a^{k-l} b^{l}$$
$$\sum_{m=0}^{N} P(m,N,p) = \sum_{m=0}^{N} {\binom{N}{m}} p^{m} q^{N-m} = (p+q)^{N} = 1$$

- binomial distribution is properly normalized
- Mean of binomial distribution:

$$\mu = \frac{\sum_{m=0}^{N} mP(m, N, p)}{\sum_{m=0}^{N} P(m, N, p)} = \sum_{m=0}^{N} mP(m, N, p) = \sum_{m=0}^{N} m\binom{N}{m} p^{m} q^{N-m}$$

 $\star$  A cute way of evaluating the above sum is to take the derivative:

$$\frac{\partial}{\partial p} \left[ \sum_{m=0}^{N} \binom{N}{m} p^{m} q^{N-m} \right] = 0$$

$$\sum_{m=0}^{N} m\binom{N}{m} p^{m-1} q^{N-m} - \sum_{m=0}^{N} \binom{N}{m} p^{m} (N-m)(1-p)^{N-m-1} = 0$$

$$p^{-1} \sum_{m=0}^{N} m\binom{N}{m} p^{m} q^{N-m} = N(1-p)^{-1} \sum_{m=0}^{N} \binom{N}{m} p^{m} (1-p)^{N-m} - (1-p)^{-1} \sum_{m=0}^{N} m\binom{N}{m} p^{m} (1-p)^{N-m}$$

$$p^{-1} \mu = N(1-p)^{-1} \cdot 1 - (1-p)^{-1} \mu$$

$$\mu = Np$$

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• Variance of binomial distribution (obtained using similar trick):

$$\sigma^{2} = \frac{\sum_{m=0}^{N} (m-\mu)^{2} P(m,N,p)}{\sum_{m=0}^{N} P(m,N,p)} = Npq$$

- $\star$  Example: Suppose you observed *m* special events (success) in a sample of *N* events
  - measured probability ("efficiency") for a special event to occur:

$$\varepsilon = \frac{m}{N}$$

error on the probability ("error on the efficiency"):

$$\sigma_{\varepsilon} = \frac{\sigma_m}{N} = \frac{\sqrt{Npq}}{N} = \frac{\sqrt{N\varepsilon(1-\varepsilon)}}{N} = \sqrt{\frac{\varepsilon(1-\varepsilon)}{N}}$$

sample (*N*) should be as large as possible to reduce uncertainty in the probability measurement

- ★ Example: Suppose a baseball player's batting average is 0.33 (1 for 3 on average).
  - Consider the case where the player either gets a hit or makes an out (forget about walks here!). probability for a hit: p = 0.33probability for "no hit": q = 1 - p = 0.67
  - On average how many hits does the player get in 100 at bats?

 $\mu = Np = 100.0.33 = 33$  hits

• What's the standard deviation for the number of hits in 100 at bats?

$$\sigma = (Npq)^{1/2} = (100 \cdot 0.33 \cdot 0.67)^{1/2} \approx 4.7$$
 hits

we expect  $\approx 33 \pm 5$  hits per 100 at bats

## **Poisson Probability Distribution**

- A widely used discrete probability distribution
- Consider the following conditions:
  - $\bullet$  p is very small and approaches 0
    - example: a 100 sided dice instead of a 6 sided dice, p = 1/100 instead of 1/6
    - example: a 1000 sided dice, p = 1/1000
  - *N* is very large and approaches  $\infty$
  - example: throwing 100 or 1000 dice instead of 2 dice
  - product Np is finite
- Example: radioactive decay
  - ★ Suppose we have 25 mg of an element
    - very large number of atoms:  $N \approx 10^{20}$
  - \* Suppose the lifetime of this element  $\lambda = 10^{12}$  years  $\approx 5 \times 10^{19}$  seconds
    - reprobability of a given nucleus to decay in one second is very small:  $p = 1/\lambda = 2x10^{-20}/\text{sec}$
    - $\square$  Np = 2/sec finite!
    - number of counts in a time interval is a Poisson process
- Poisson distribution can be derived by taking the appropriate limits of the binomial distribution

$$\begin{aligned} P(m,N,p) &= \frac{N!}{m!(N-m)!} p^m q^{N-m} \\ \frac{N!}{(N-m)!} &= \frac{N(N-1)\cdots(N-m+1)(N-m)!}{(N-m)!} = N^m \\ q^{N-m} &= (1-p)^{N-m} = 1 - p(N-m) + \frac{p^2(N-m)(N-m-1)}{2!} + \dots \approx 1 - pN + \frac{(pN)^2}{2!} + \dots \approx e^{-pN} \end{aligned}$$

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Siméon Denis Poisson June 21, 1781-April 25, 1840

$$P(m,N,p) = \frac{N^m}{m!} p^m e^{-pN}$$
Let  $\mu = Np$ 

$$P(m,\mu) = \frac{e^{-\mu}\mu^m}{m!}$$

$$\sum_{m=0}^{m=\infty} \frac{e^{-\mu}\mu^m}{m!} = e^{-\mu}\sum_{m=0}^{m=\infty} \frac{\mu^m}{m!} = e^{-\mu}e^{\mu} = 1$$
Poisson distribution is normalized
• *m* is always an integer > 0
•  $\mu$  does not have to be an integer
• It is easy to show that:
$$\mu = Np = \text{mean of a Poisson distribution}$$
mean and variance are the same number
 $\sigma^2 = Np = \mu$  = variance of a Poisson distribution
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• Radioactivity example with an average of 2 decays/sec:
• What's the probability of zero decays in one second?
$$p(0,2) = \frac{e^{-2}2^0}{0!} = \frac{e^{-2} \cdot 1}{1} = e^{-2} = 0.135 \rightarrow 13.5\%$$
• What's the probability of more than one decay in one second?
$$p(>1,2) = 1 - p(0,2) - p(1,2) = 1 - \frac{e^{-2}2^0}{0!} - \frac{e^{-2}2^1}{1!} = 1 - e^{-2} - 2e^{-2} = 0.594 \rightarrow 59.4\%$$
• Estimate the most probable number of decays/sec?
$$\frac{\partial}{\partial m} P(m,\mu)\Big|_{m^*} = 0$$

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• To solve this problem its convenient to maximize  $\ln P(m, \mu)$  instead of  $P(m, \mu)$ .

$$\ln P(m,\mu) = \ln \left(\frac{e^{-\mu}\mu^m}{m!}\right) = -\mu + m \ln \mu - \ln m!$$

• In order to handle the factorial when take the derivative we use *Stirling's Approximation*:  $\ln m! \approx m \ln m - m$ 

$$\frac{\partial}{\partial m} \ln P(m,\mu) = \frac{\partial}{\partial m} (-\mu + m \ln \mu - \ln m!)$$
$$\approx \frac{\partial}{\partial m} (-\mu + m \ln \mu - m \ln m + m)$$
$$= \ln \mu - \ln m - m \frac{1}{m} + 1$$
$$= 0$$
$$m^* = \mu$$

- The most probable value for *m* is just the average of the distribution
- If you observed *m* events in an experiment, the error on *m* is  $\sigma = \sqrt{\mu} = \sqrt{m}$
- •\_This is only approximate since Stirlings Approximation is only valid for large *m*.
- Strictly speaking m can only take on integer values while  $\mu$  is not restricted to be an integer.

## Comparison of Binomial and Poisson distributions with mean $\mu = 1$

