# Statistical Consistency of Multipartite Ranking

Yoonkyung Lee\*1
Department of Statistics
The Ohio State University
\*joint work with Kazuki Uematsu

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## Ranking

- Aims to order a set of objects or instances reflecting their underlying utility, relevance or quality.
- Has gained increasing attention in machine learning, collaborative filtering and information retrieval for website search and document retrieval.



(Source: Google images of "ranking")



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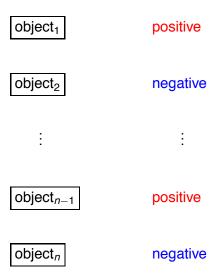
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# Data for Ranking



How to order objects so that positive cases are ranked higher than negative cases?

### Main Questions

- ▶ How to rank?
- What loss criteria to use for ranking?
- What is the best ranking function given a criterion?
- How is it related to the underlying probability distribution for data?
- How to learn a ranking function from data?

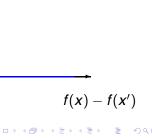
### **Notation**

- ▶  $X \in \mathcal{X}$ : an instance to rank
- ▶  $Y \in \mathcal{Y} = \{1, \dots, k\}$ : an ordinal response in multipartite ranking (bipartite ranking when k = 2)
- ▶  $f: \mathcal{X} \to \mathbb{R}$ : a real-valued ranking function whose scores induce ordering over the input space
- ▶ Training data: n pairs of (X, Y) from  $\mathcal{X} \times \mathcal{Y}$

## Pairwise Ranking Loss

For a pair of "positive" x and "negative" x', define a loss of ranking function f as

$$\ell_0(f;x,x') = I(f(x) - f(x') < 0) + \frac{1}{2}I(f(x) - f(x') = 0)$$



# Bipartite Ranking

- Note the invariance of the pairwise loss under order-preserving transformations.
- Find f minimizing the empirical ranking risk

$$R_{n_+,n_-}(f) = \frac{1}{n_+ n_-} \sum_{i=1}^{n_+} \sum_{j=1}^{n_-} \ell_0(f; \mathbf{x}_i, \mathbf{x}_j')$$

- Minimizing ranking error is equivalent to maximizing AUC (area under ROC curve).
- Use convex surrogate loss for risk minimization.
   (e.g. RankBoost, RankSVM, and RankNet)

# Likelihood Ratio Minimizes Ranking Risk

Clémençon et al. (2008) and Uematsu and Lee (2011)

### **Theorem**

Define  $f_0^*(x) = g_+(x)/g_-(x)$ , and let  $R_0(f) = E(\ell_0(f; X, X'))$  denote the ranking risk of f under the bipartite ranking loss. Then for any ranking function f,

$$R_0(f_0^*) \leq R_0(f).$$

# Extension to Multipartite Ranking

▶ In general  $(k \ge 2)$ , for a pair of (x, y) and (x', y') with y > y', define a loss of ranking function f as

$$\ell_0(f; x, x', y, y') = c_{y'y}I(f(x) < f(x')) + \frac{1}{2}c_{y'y}I(f(x) = f(x'))$$

where  $c_{y'y}$  is the cost of misranking a pair of y and y'. (Waegeman et al. 2008)

► Again, \(\ell\_0\) is invariant under order-preserving transformations.

# Optimal Ranking Function for Multipartite Ranking

#### **Theorem**

(i) When k = 3, let  $f_0^*(x) = \frac{c_{12}P(Y=2|x) + c_{13}P(Y=3|x)}{c_{13}P(Y=1|x) + c_{23}P(Y=2|x)}$ . Then for any ranking function f,

$$R_0(f_0^*; \mathbf{c}) \leq R_0(f; \mathbf{c}).$$

(ii) When 
$$k > 3$$
 and let  $f_0^*(x) = \frac{\sum_{i=2}^k c_{1i} P(Y=i|x)}{\sum_{j=1}^{k-1} c_{jK} P(Y=j|x)}$ . If  $c_{1k}c_{ji} = c_{1i}c_{jk} - c_{1j}c_{ik}$  for all  $1 < j < i < k$ , then for any ranking function  $f$ ,

$$R_0(f_0^*; \mathbf{c}) \leq R_0(f; \mathbf{c}).$$

## Remark

Let  $c_{ji} = (s_i - s_j)w_iw_j I(i > j)$  for some increasing scale  $\{s_j\}_{j=1}^k$  and non-negative weight  $\{w_j\}_{i=1}^k$ . e.g.  $c_{ji} = (i - j)I(i \ge j)$ 

# **Ordinal Regression**

Ordinal regression is commonly used to analyze data with ordinal response in practice.

$$y = 1 \qquad 2 \qquad \cdots \qquad k-1 \qquad k$$

$$-\infty = \theta_0 \qquad \theta_1 \qquad \theta_2 \qquad \cdots \qquad \theta_{k-2} \qquad \theta_{k-1} \qquad \theta_k = \infty$$

▶ A typical form of loss in ordinal regression for f with thresholds  $\{\theta_j\}_{j=1}^{k-1}$ :

$$\ell(f, \{\theta_j\}_{j=1}^{k-1}; x, y) = \ell(f(x) - \theta_{y-1}) + \ell(\theta_y - f(x)),$$

where  $\theta_0 = -\infty$  and  $\theta_k = \infty$ .



# Convex Loss in Ordinal Regression

ORBoost (Lin and Li 2006):

$$\ell(s) = \exp(-s)$$

Proportional Odds model (McCullagh 1980, Rennie 2006):

$$\ell(s) = \log(1 + \exp(-s))$$

Support Vector Ordinal Regression (Herbrich et al. 2000):

$$\ell(s) = (1-s)_+$$

# Ordinal Regression Boosting (ORBoost)

▶ The optimal ranking function  $f^*$  under  $\ell(s) = \exp(-s)$  is

$$f^*(x) = \frac{1}{2} \log \frac{\sum_{i=2}^k P(Y = i | x) \exp(\theta_{i-1}^*)}{\sum_{j=1}^{k-1} P(Y = j | x) \exp(-\theta_j^*)}$$

where  $\theta_j^*$  are constants depending only on  $P_{X,Y}$ .

▶ When *k* = 3,

$$f^*(x) = \frac{1}{2} \log \frac{P(Y=2|x) + \exp(\theta_2^* - \theta_1^*) P(Y=3|x)}{\exp(\theta_2^* - \theta_1^*) P(Y=1|x) + P(Y=2|x)}$$

up to a constant. Hence,  $f^*$  preserves the ordering of  $f_0^*$  with  $c_{12}=c_{23}=1$  and  $c_{13}=e^{\theta_2^*-\theta_1^*}$ .

## **Proportional Odds Model**

Cumulative logits (*McCullagh* 1980)

$$\log \frac{P(Y \leq j|x)}{P(Y > j|x)} = f(x) - \theta_j,$$

where 
$$-\infty = \theta_0 < \theta_1 < \ldots < \theta_{k-1} < \theta_k = \infty$$
.

- ▶ Given  $\{\theta_j\}_{j=1}^{k-1}$ , maximizing the log likelihood amounts to ordinal regression with  $\ell(s) = \log(1 + \exp(-s))$ .
- ▶ When k = 3, given  $\theta_1$  and  $\theta_2$ , the minimizer of the deviance risk  $f^*$  satisfies

$$\exp(f^*(x)) = \frac{q(x) - 1 + \sqrt{(q(x) - 1)^2 + 4\exp(\theta_1 - \theta_2)q(x)}}{2\exp(-\theta_2)},$$

where 
$$q(x) = \frac{P(Y = 2|x) + P(Y = 3|x)}{P(Y = 1|x) + P(Y = 2|x)} = f_0^*(x)$$
 with  $c_{12} = c_{23} = c_{13} = 1$ .

▶ When  $\theta_2 > \theta_1$ ,  $f^*(x)$  preserves the ordering of q(x).

# Support Vector Ordinal Regression

SVOR with Implicit constraints in Chu and Keerthi (2007)

$$\ell(r, \{\theta_j\}_{j=1}^{k-1}; x, y) = \sum_{j=1}^{y-1} (1 - (f(x) - \theta_j))_+ + \sum_{j=y}^{k-1} (1 - (\theta_j - f(x)))_+.$$

▶ When k = 3,  $f^*(x)$  is a step function of  $r(x) = \frac{p_2(x) + p_3(x)}{p_1(x) + p_2(x)}$  (i.e.  $f_0^*$  with  $c_{12} = c_{13} = c_{23}$ ).

$$r(x)$$
  $(0,\frac{1}{2})$   $(\frac{1}{2},1)$   $(1,2)$   $(2,\infty)$   $f^*(x)$   $\theta_1-1$   $\min(\theta_1+1,\theta_2-1)$   $\max(\theta_1+1,\theta_2-1)$   $\theta_2+1$ 

### Numerical Illustration

Simulation setting:

$$X|Y = 1 \sim N(-2, 1), X|Y = 2 \sim N(0, 1)$$
 and  $X|Y = 3 \sim N(2, 1)$ 

• When  $c_{12} = c_{23} = c_{13} = 1$ ,

$$f_0^*(x) = \frac{P(Y=2|X=x) + P(Y=3|X=x)}{P(Y=1|X=x) + P(Y=2|X=x)} = \frac{e^{2x} + e^2}{e^{-2x} + e^2}.$$

- Generate 500 observations in each category.
- Apply pairwise ranking risk minimization with exponential loss, proportional odds model, ORBoost and SVOR.

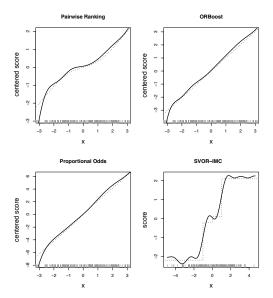


Figure: Theoretical ranking function (dotted line) and estimated ranking function (solid line) for pairwise ranking risk minimization with exponential loss, ORBoost, proportional odds model and SVOR with implicit constraints.

# Application to Movie-Lens Data

- ► The data set consists of 100,000 ratings (on a scale of 1 to 5) for 1,682 movies by 943 users (GroupLens-Research).
- Contains content information about the movies (release date and genres) and demographic information about the users (age, gender and occupation).
- Transform five categories into three categories:
   "Low" (1-3), "Middle" (4) and "High" (5)
   and check the analytical results in k = 3.

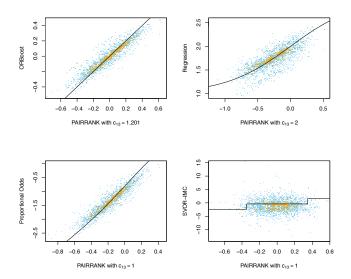


Figure: Scatter plots of ranking scores from ORBoost, regression, proportional odds model, and SVOR against pairwise ranking scores with matching cost  $c_{13}$  for MovieLens data with three categories. The solid lines indicate theoretical relation between ranking scores.

# Effect of Differential Ranking Cost

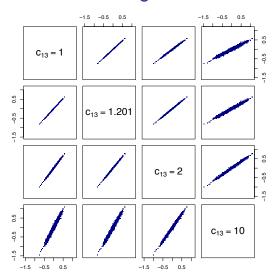


Figure: Scatter plots of pairwise ranking scores (centered to zero) with different ranking cost  $c_{13}$  for MovieLens data when  $c_{12} = c_{23} = 1$ .



# **Concluding Remarks**

- Provide a statistical view of ranking by identifying the optimal ranking function given loss criteria
- For pairwise multipartite ranking, the optimal ranking depends on the ratio of conditional probability weighted by misranking costs.
- The solution to some ordinal regression methods can be viewed as a special case of the optimal function in multipartite ranking.
- Our study bridges traditional methods such as proportional odds model in statistics with ranking algorithms in machine learning.

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